A New Interpolative Perspective on Shape-restricted Estimation

Kenta Takatsu, Carnegie Mellon University Joint work with Tianyu Zhang and Arun Kuchibhotla WNAR June, 2025 From Isotonic to Lipschitz Regression:

A New Interpolative Perspective on Shape-restricted Estimation

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Abstract

This manuscript seeks to bridge two seemingly disjoint paradigms of nonparametric regression: estimation based on smoothness assumptions and shape constraints. The proposed approach is motivated by a conceptually simple observation: Every Lipschitz function is a sum of monotonic and linear functions. This principle is further generalized to the higher-order monotonicity and multivariate covariates. A family of estimators is proposed based on a sample-splitting procedure, which inherits desirable methodological, theoretical, and computational properties of shape-restricted estimators. The theoretical analysis provides convergence guarantees of the estimator under heteroscedastic and heavy-tailed errors, as well as adaptivity properties to the unknown complexity of the true regression function. The generality of the proposed decomposition framework is demonstrated through new approximation results, and extensive numerical studies validate the theoretical properties of the proposed estimation framework.

Keywords— Nonparametric regression, Model selection, Shape-restricted estimation, Constructive Approximation, Heavy-tailed data

arXiv:2307.05732

Observe IID real-valued R.V.
$$(X_1, Y_1), \ldots, (X_n, Y_n) \in [0,1] \times \mathbb{R}$$
 such that $Y_i = f_0(X_i) + \varepsilon_i$ where $f_0(x) := \mathbb{E}[Y | X = x]$.

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Choice 2 Assume that f_0 has some shape (e.g., monotonic).

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Without additional structure, the minimax rates are $||\hat{f} - f_0||^2 = O_P(n^{-2/3})$ for both choices.

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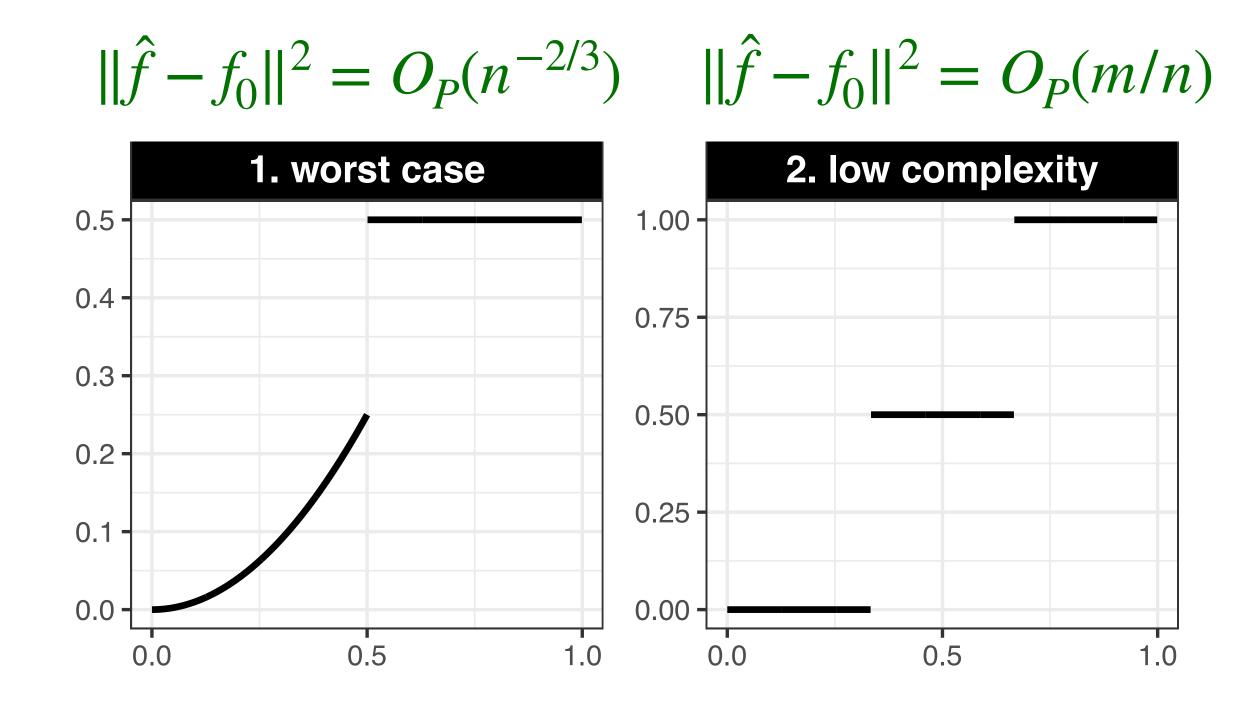
Choice 2 Assume that f_0 has some shape (e.g., monotonic).

Without additional structure, the minimax rates are $||\hat{f} - f_0||^2 = O_P(n^{-2/3})$ for both choices.

Q. Are there connections between the spaces of monotonic and once-differentiable functions?

Properties of Monotone Estimator

- 1. Many shape-restricted estimators (e.g., LSEs) are tuning parameter free.
- 2. They converge at adaptive rates;



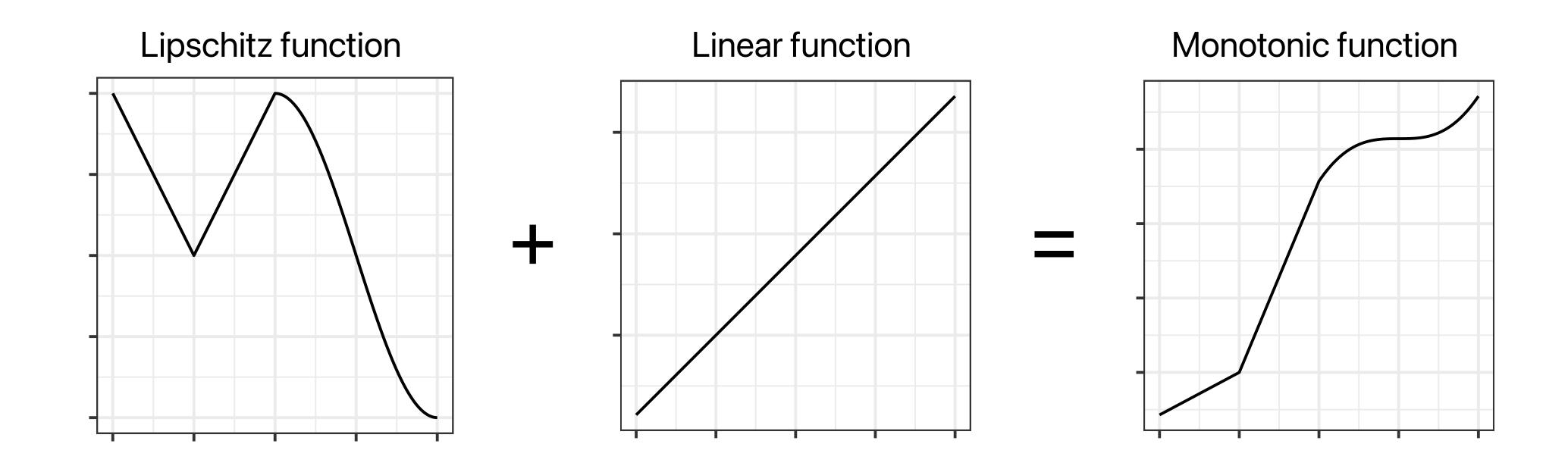
Q. Can we identify structural links between smooth and monotone functions?

Q. Can we construct "tuning parameter free" and adaptive estimators beyond shape-restricted problems?

Q. Can we develop an optimal estimator for both smooth and shape-restricted classes simultaneously?

Proposition 3.1

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Adding a linear term to a monotone estimator significantly increases its expressiveness.

We introduce a new decomposition class:

$$\mathcal{F}_1(L) := \big\{ f \colon [0,1] \mapsto \mathbb{R} \ : f(x) = g(x) - Lx, \exists g \text{ is non-decr} \big\}.$$

The class $\mathcal{F}_1(0)$ is equivalent to the space of monotone functions. As $k \uparrow L$, the space $\mathcal{F}_1(k)$ becomes larger and eventually includes all L-Lipschitz functions (and more).

Proposition 3.1

A Lipschitz estimator can take the form: $\hat{f}(x) = \hat{g}(x) - \hat{L}x$.

Lipschitz est.

 \Rightarrow

Monotone est.

+

Model selection

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A Lipschitz estimator can take the form: $\hat{f}(x) = \hat{g}(x) - \hat{L}x$.

Lipschitz est. \Rightarrow Monotone est. + Model selection

Use sample-splitting to avoid overfitting.

This leads to two stages of least-squares.

Estimation Procedure

1. Split data into two parts D_1 and D_2 . Prepare a set \mathscr{L} for candidate L (e.g., $\mathscr{L} = [-\log n, \log n]$).

In practice, you do not need to specify \mathscr{L} .

The optimization program (e.g., optim(\cdot) in R) will handle this.

Estimation Procedure

- 1. Split data into two parts D_1 and D_2 . Prepare a set \mathscr{L} for candidate L (e.g., $\mathscr{L} = [-\log n, \log n]$).
- 2. For each $L \in \mathcal{L}$, compute isotonic regression (monotone LSE) on D_1 :

$$\widehat{g}_L := \underset{g:\text{monotone}}{\text{arg min}} \sum_{(x,y) \in D_1} ((y + Lx) - g(x))^2 \text{ and return } \widehat{f}_L(x) := \widehat{g}_L(x) - Lx.$$

Step 2 can be computed using near linear time algorithm (e.g., isoreg(\cdot) in R).

Estimation Procedure

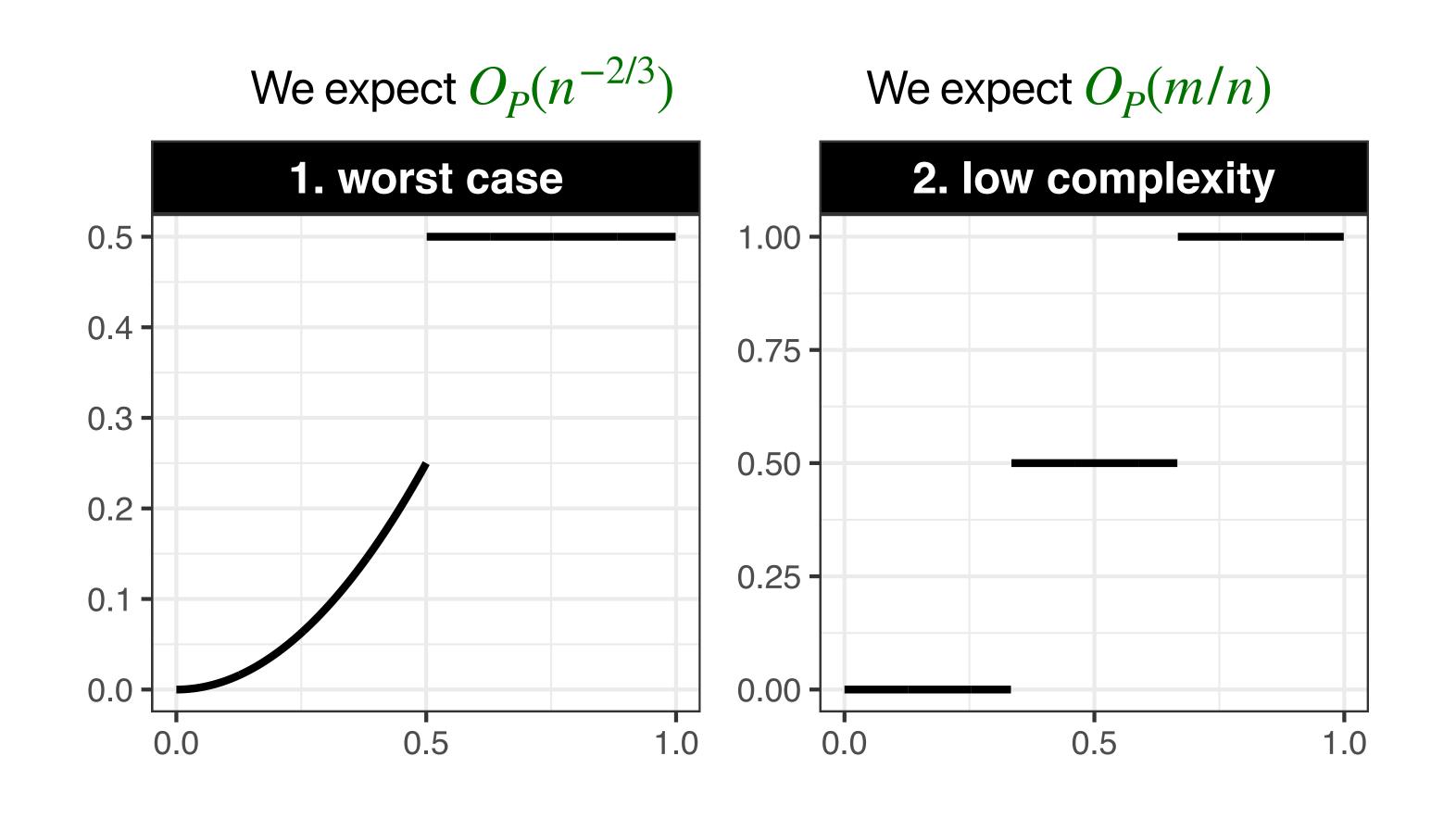
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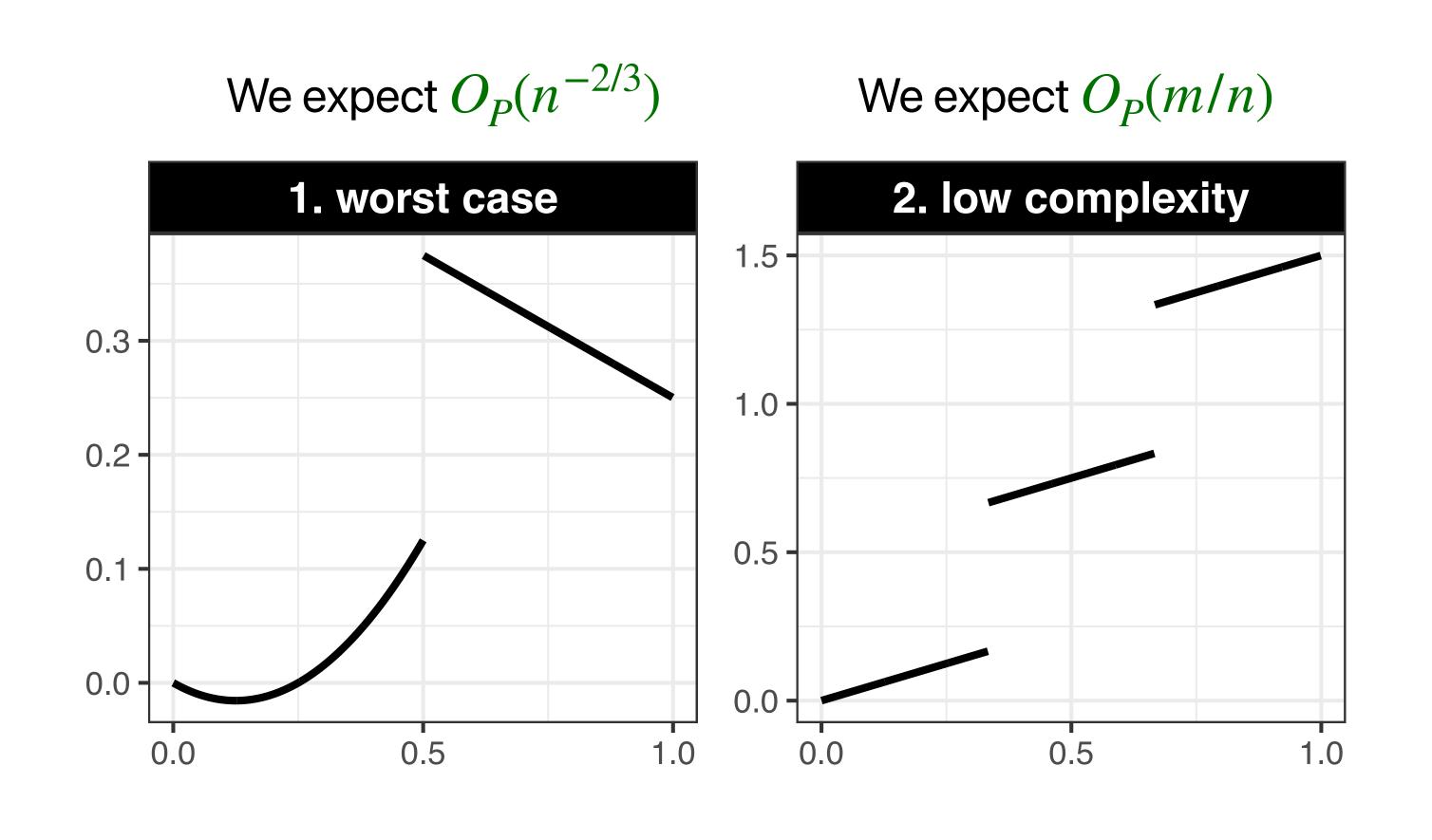
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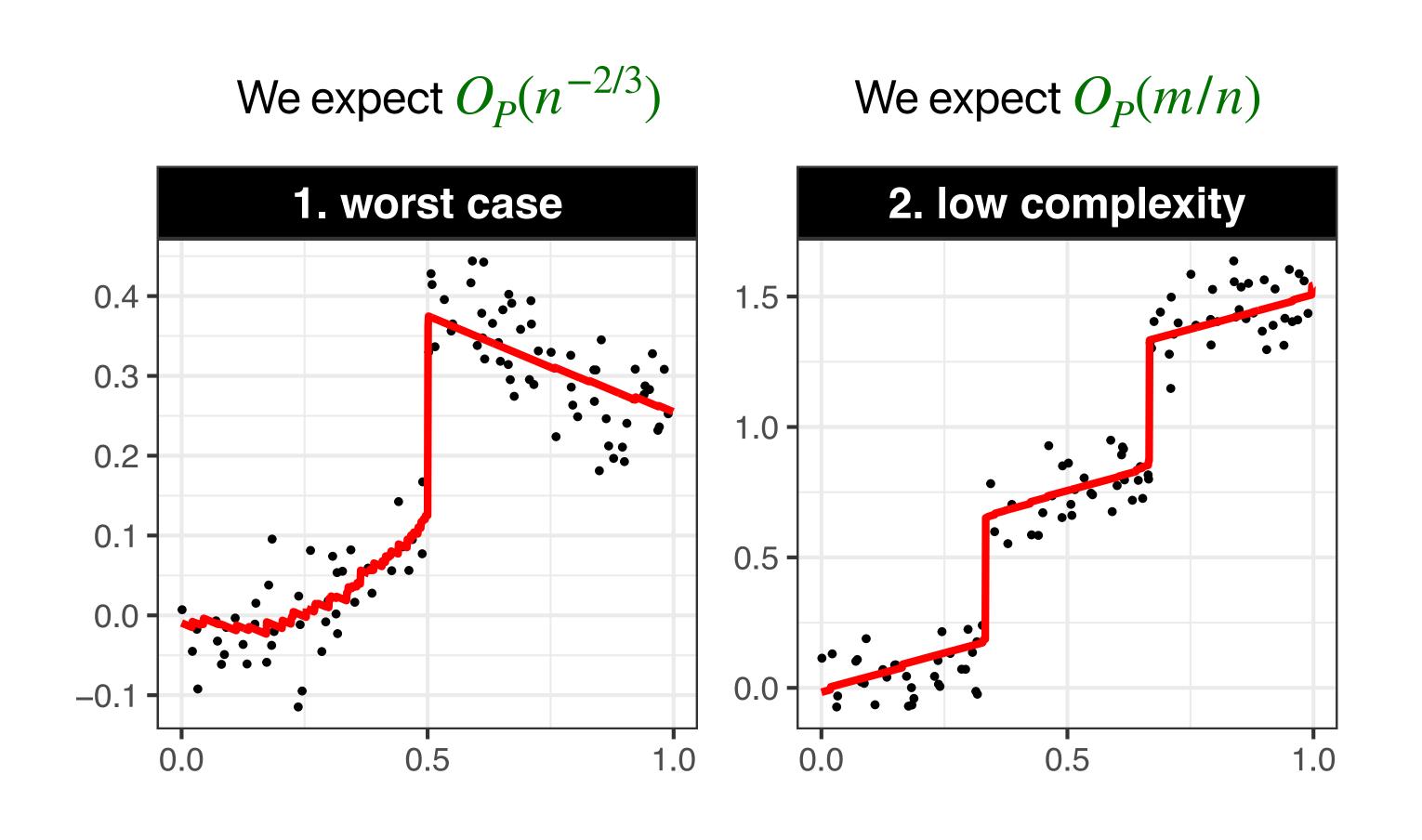
3. Perform model selection over \mathscr{L} on D_2 , for instance,

$$\widehat{L} := \underset{L \in \mathcal{L}}{\operatorname{arg\,min}} \sum_{(x,y) \in D_2} (y - \widehat{f}_L(x))^2 \text{ and return } \widehat{f}(x) := \widehat{g}_{\widehat{L}}(x) - \widehat{L}x.$$

One can use other "robust" cross-validation using the median-of-means or aggregation.







Lipschitz est. \Rightarrow Monotone est. + Model Selection

For each fixed $L \in \mathcal{L}$, fitting $y + Lx \sim x$ using monotone regression \widehat{g}_L is (almost) identical to the standard estimation for shape-restricted problems.

Selecting the "best" $L \in \mathcal{L}$ from a set $\{L \mapsto \widehat{g}_L(x) - Lx\}$ is essentially a model selection performed over a nonparametric function space.

$$\|\hat{f} - f_0\|^2 = O_P\left(\inf_{L \in \mathcal{L}} R_{n,L}^{\text{mono}} + R_{n,\mathcal{L}}^{\text{CV}}\right)$$

Recall
$$Y_i = f_0(X_i) + \varepsilon_i$$

- ε_i has finite q moments conditioning on X_i i.e., $\mathbb{E}[\,|\,\varepsilon_i\,|^q\,|\,X_i] \leq C$.
- **A2** $ε_i$ is β-sub-Weibull conditioning on X_i (β = 1 corresponds to sub-exponential and β = 2 corresponds to sub-Gaussian).

Worst case Low complexity

A1

Theorem 4.1-4.3

Recall
$$Y_i = f_0(X_i) + \varepsilon_i$$

- ε_i has finite q moments conditioning on X_i i.e., $\mathbb{E}[|\varepsilon_i|^q|X_i] \leq C$. **A1**
- ε_i is β -sub-Weibull conditioning on X_i ($\beta=1$ corresponds to sub-**A2** exponential and $\beta = 2$ corresponds to sub-Gaussian).

Theorem 4.1-4.3

Low complexity

A2
$$O(n^{-2/3}) + O(n^{-1})$$
 $O(m/n) + O(n^{-1})$

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$$R_{n,L}^{\text{mono}} + R_{n,L}^{\text{CV}}$$

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A1 $O(n^{-2/3}) + O(n^{-1+1/q})$ $O(m/n) + O(n^{-1+1/q})$ A2 $O(n^{-2/3}) + O(n^{-1})$ $O(m/n) + O(n^{-1})$ $O(m/n) + O(n^{-1})$ $O(m/n) + O(n^{-1})$

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Theorem 4.1-4.3

 $R_{n,L}^{\text{mono}} + R_{n,L}^{\text{CV}}$

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Lipschitz est. \Rightarrow Monotone est. + Model selection

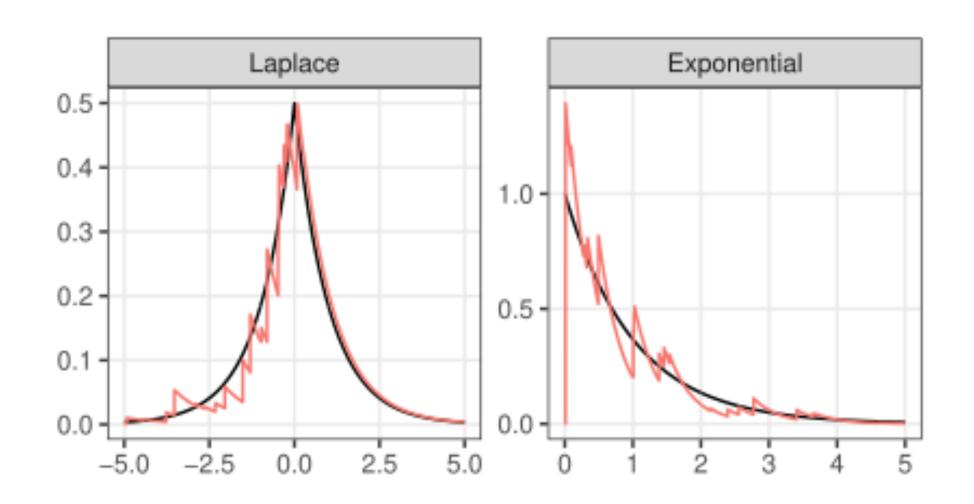
Under finite q moments, cross-validation becomes slower than estimation. This is because cross-validation based on squared error is *suboptimal* for heavy-tailed problems. We show that model selection based on median-of-means achieves the optimal rate under both **A1** and **A2**.

Open Problem: Irregular Density Estimation

Observe $X_1, ..., X_n$ from a density function f_0 such that $\log f_0(x) = g(x) - Lx$ where g is monotone.

This problem can be considered as *log Lipschitz estimation* in compaction to log concave estimation.

The example of the worst case is Laplace distribution and the adaptive case is exponential distribution.



Open Problem: Shape Detection

Suppose that the estimator is obtained by the proposed method: \hat{r}

$$\hat{f}(x) = \hat{g}(x) - \hat{L}x.$$

Intuitively, $|\hat{L}| \approx 0$ should indicate that g_0 is monotone while $|\hat{L}| \gg 0$ can be the evidence against monotonicity.

(Caveat) The population L_0 is not identifiable.

Summary

A nonparametric regression can be decomposed into two problems: (1) shape-restricted estimation and (2) cross-validation.

We propose a new nonparametric estimator that converges at adaptive rates under the some assumption on the error ε_i (e.g., sub-Weibull).

When the error ε_i has finite q moments, cross-validation can be "harder" than estimation. This cannot be improved without considering an alternative model selection procedure.

In the manuscript, we discuss some "higher-order" (e.g., convex + quadratic) and multivariate extensions.

Higher-order Methods



Higher-order Methods



Thank you

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This manuscript seeks to bridge two seemingly disjoint paradigms of nonparametric regression: estimation based on smoothness assumptions and shape constraints. The proposed approach is motivated by a conceptually simple observation: Every Lipschitz function is a sum of monotonic and linear functions. This principle is further generalized to the higher-order monotonicity and multivariate covariates. A family of estimators is proposed based on a sample-splitting procedure, which inherits desirable methodological, theoretical, and computational properties of shape-restricted estimators. The theoretical analysis provides convergence guarantees of the estimator under heteroscedastic and heavy-tailed errors, as well as adaptivity properties to the unknown complexity of the true regression function. The generality of the proposed decomposition framework is demonstrated through new approximation results, and extensive numerical studies validate the theoretical properties of the proposed estimation framework.

Keywords— Nonparametric regression, Model selection, Shape-restricted estimation, Constructive Approximation, Heavy-tailed data

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Appendix

Oracle Inequality for Model Selection

We develop a new general oracle inequality regarding an LSE:

Let \hat{f} be an LSE over an arbitrary uniformly bounded and finite class \mathscr{F} , and ε_i has finite conditional q moments, then for any $\delta>0$

$$\mathbb{E}\|\hat{f} - f_0\|^2 \le (1 + \delta) \inf_{f \in \mathcal{F}} \|f - f_0\|^2 + \left(n^{-1 + 1/q} + \frac{1}{\delta n}\right) \log(|\mathcal{F}|)$$

Theorem 4.4

The issue is the term $n^{-1+1/q}$, and this is *not* generally improvable for LSEs (Han and Wellner (2018); Kuchibhotla and Patra (2022)).

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Mourtada, J. (2022). Exact minimax risk for linear least squares, and the lower tail of sample covariance matrices. *The Annals of Statistics*, 50(4):2157–2178.

Thank You

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Takatsu, K. and Kuchibhotla, A. K. (2025). Bridging Root-n and Non-standard Asymptotics: Dimension-agnostic Adaptive Inference in M-Estimation, arXiv:2501.07772.

Takatsu, K. and Kuchibhotla, A. K. (2024). Generalized van Trees Inequality: Local Minimax Bounds for Non-differentiable Functionals and Irregular Statistical Models. arXiv:2405.06437.

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$$Y_i = f_0(X_i) + \varepsilon_i$$

(A1) ε_i has finite q moments conditioning on X_i i.e., $\mathbb{E}[\,|\,\varepsilon_i\,|^q\,|\,X_i] \leq C$

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Recall $\mathcal{F}_1(L):=\{f\colon [0,1]\mapsto \mathbb{R}\,:\, g(x)-Lx, \exists g \text{ is non-decr}\}$ and let \mathscr{L} be a finite set of linear parameter L

Theorem 4.1-4.3 (informal)

The following "two-layer" oracle inequality holds for any $\varepsilon \in (0,1)$

$$\|\hat{f} - f_0\|^2 \le \inf_{L \in \mathcal{L}} \left\{ C \inf_{f \in \mathcal{F}_1(L)} \|f - f_0\|^2 + (R_{n,L}^{\text{mono}} + \log(|\mathcal{L}|)R_{n,L}^{\text{CV}}) \right\}$$

Proposition 3.1

Define *L*-Lipschitz class as:

$$\Sigma_1(L) := \left\{ f : [0,1] \mapsto \mathbb{R} : \frac{|f(x_1) - f(x_2)|}{|x_1 - x_2|} \le L, \forall x_1, x_2 \in [0,1] \right\}.$$

We introduce a new decomposition class:

$$\mathcal{F}_1(L) := \big\{ f \colon [0,1] \mapsto \mathbb{R} \ : f(x) = g(x) - Lx, \exists g \text{ is non-decr} \big\}.$$

We can show that $\Sigma_1(L) \subsetneq \mathscr{F}_1(L)$ and $\Sigma_1(L)$ is not dense in $\mathscr{F}_1(L)$ for fixed L.

New Function Spaces

For an integer $r \ge 1$ and h satisfying $h \in [0, 1 - rh]$, we define a forward operator:

$$\Delta_h^r(f, x) := \sum_{m=0}^r \binom{r}{m} (-1)^{r-m} f(x + mh).$$

Let $\mathscr{C}(k)$ be the collection of k-monotone functions (Chatterjee et al., 2015):

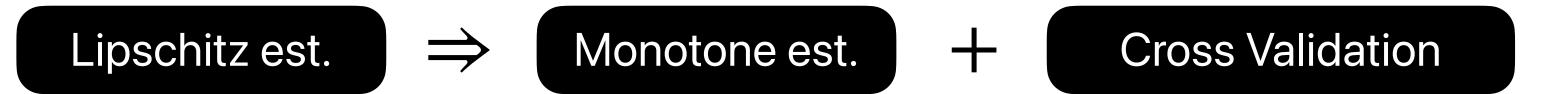
$$\mathscr{C}(k) := \left\{ g : [0,1] \to \mathbb{R} : \Delta_h^k(g,x) \ge 0 \text{ for all } x \in [0,1] \right\}$$

Examples: $\mathscr{C}(1) := \{ \text{non-decr. fns.} \}$ and $\mathscr{C}(2) := \{ \text{convex fns.} \}$

We introduce a new function space:

$$\mathscr{F}(k,L) := \{g(x) - (L/k!)x^k \text{ such that } g \in \mathscr{C}(k)\}.$$

Properties of the Estimator



(Intuition) When monotone estimation is *harder* than cross-validation, the convergence rate of the proposed estimator should match that of a monotone estimator.

We introduce a new function space:

$$\mathscr{F}(k,L) := \{g(x) - (L/k!)x^k \text{ such that } g \in \mathscr{C}(k)\}.$$

Define kth bounded L-Lipschitz class for an integer k as:

$$\Sigma_k(L) := \left\{ f : [0,1] \mapsto \mathbb{R} : \sum_{0 \le m \le k-1} \|D^m f\|_{\infty} + \frac{|D^{k-1} f(x_1) - D^{k-1} f(x_2)|}{|x_1 - x_2|} \le L, \forall x_1, x_2 \in [0,1] \right\}$$

We can show that $\Sigma_k(L) \subsetneq \mathscr{F}(k,L)$ and $\Sigma_k(L)$ is not dense in $\mathscr{F}(k,L)$

^{*} $D^k f$ denotes the kth weak derivative and $D^0 f = f$